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The Existence of Maximum Likelihood Estimates for the Binary Response Logistic Regression Model

William F. McCarthy

Abstract

The existence of maximum likelihood estimates for the binary response logistic regression model depends on the configuration of the data points in your data set. There are three mutually exclusive and exhaustive categories for the configuration of data points in a data set: Complete Separation, Quasi-Complete Separation, and Overlap. For this paper, a binary response logistic regression model is considered. A 2 x 2 tabular presentation of the data set to be modeled is provided for each of the three categories mentioned above. In addition, the paper will present an example of a data set whose data points have a linear dependency. Both unconditional maximum likelihood estimation (asymptotic inference) and exact conditional estimation (exact inference) will be considered and contrasted in terms of results. The statistical software package SAS will be used for the binary response logistic regression modeling.

Introduction

The Existence of Maximum Likelihood Estimates for the Logistic Regression Model depends on the configuration of the data points in your data set (Albert and Anderson, 1984; Santner and Duffy, 1985; So, 1995). There are three mutually exclusive and exhaustive categories for the configuration of data points in a data set:

- Complete Separation
- Quasi-Complete Separation
- Overlap

Refer to So (1995) for a nice graphical illustration of these three categories. For this paper, a binary response logistic regression model is considered. A 2 x 2 tabular presentation of the data set to be modeled is provided for each of the three categories mentioned above. In addition, the paper will present an example of a data set whose data points have a linear dependency.

Unconditional maximum likelihood estimation (asymptotic inference) is used when matched data are not considered, provided that the total number of variables in the model is not too large relative to the number of observations (Kleinbaum, 1994). This method of inference is based on maximizing the likelihood function for parameter estimation using the unconditional formula (Kleinbaum, 1994). This is the usual large-sample asymptotic method used by most of the current statistical software packages (Kleinbaum, 1994; Mehta and Patel, 1995).

The existence and uniqueness of maximum likelihood parameter estimates for the logistic regression model depends on the pattern of the data points in the observation space (Albert and Anderson, 1984; Santer and Duffy, 1986; So, 1993).

Complete Separation of data points gives non-unique infinite parameter estimates. Thus, maximum likelihood parameter estimates do not exist. Quasi-Complete Separation of data points also gives non-unique infinite parameter estimates. Thus, maximum likelihood parameter estimates do not exist. Maximum likelihood parameter estimates exist and are unique when there is an Overlap of data points. Complete separation and quasi-complete separation of data points usually occur with small data sets. Complete separation can occur for any type of data, but quasi-complete separation is not likely for quantitative data.

To contrast unconditional maximum likelihood estimation, exact conditional estimation will be considered as well. The theory of exact conditional logistic regression analysis (exact inference) was first proposed by Cox (1970). The computational methods employed in the statistical software package SAS (PROC LOGISTIC) are described in Hirji, Mehta, and Patel (1987), Hirji (1992), and Mehta, Patel, and Senchaudhuri (1992).

Exact conditional inference is based on generating the conditional distribution for the sufficient statistics of the parameters of interest. This distribution is called the permutation or exact conditional distribution. If the sufficient statistic of the β being estimated lies at one extreme of its range, a median unbiased estimate is reported (Hirji, Tsiatis, and Mehta 1989).

Methods

This paper will use both asymptotic and exact inference when modeling the data and will present the SAS output obtained when each of the four data sets are used for modeling. The emphasis of this paper is to show how SAS handles these four data sets when both asymptotic and exact inference is used with respect to a binary response logistic regression modeling. The outcome variable is binary (Mutation; YES, NO) and the covariate is categorical (Drug; EXPOSURE, NON-EXPOSURE). An intercept (Constant) term will be included in the model as well.

Data Sets that will be considered in this paper:

Complete Separation

	Mutation=NO	Mutation=YES	Total
Drug=NON-EXPOSURE	0	14	14
Drug=EXPOSURE	37	0	37
Total	37	14	51

Quasi-Complete Separation

	Mutation=NO	Mutation=YES	Total
Drug=NON-EXPOSURE	12	0	12
Drug=EXPOSURE	25	14	39
Total	37	14	51

Overlap

	Mutation=NO	Mutation=YES	Total
Drug=NON-EXPOSURE	9	3	12
Drug=EXPOSURE	25	14	39
Total	34	17	51

Linear Dependency

	Mutation=NO	Mutation=YES	Total
Drug=NON-EXPOSURE	0	0	0
Drug=EXPOSURE	37	14	51
Total	37	14	51



SAS Output

The SAS output for asymptotic and exact inference when considering the Complete Separation data set is presented below.

Complete Separation

	Mutation=NO	Mutation=YES	Total
Drug=NON-EXPOSURE	0	14	14
Drug=EXPOSURE	37	0	37
Total	37	14	51

The LOGISTIC Procedure

Model Information

Data Set	WORK.TEST
Response Variable	mutation
Number of Response Levels	2
Number of Observations	51
Model	binary logit
Optimization Technique	Fisher's scoring

Response Profile

Ordered Value	mutation	Total Frequency
1	1	14

Probability modeled is mutation=1.

Model Convergence Status

Complete separation of data points detected.

WARNING: The maximum likelihood estimate does not exist.

WARNING: The LOGISTIC procedure continues in spite of the above warning. Results shown are based on the last maximum likelihood iteration. Validity of the model fit is questionable.

Model Fit Statistics

	Criterion	Intercept Only	Intercept and Covariates
Collection of Biostatistics Research Archive	AIC SC -2 Log L	61.945 63.877 59.945	4.007 7.871 0.007

Testing Global Null Hypothesis: BETA=0

Test	Chi-Square	DF	Pr > ChiSq
Likelihood Ratio	59.9375	1	<.0001
Score	51.0000	1	<.0001
Wald	0.2295	1	0.6319

Analysis of Maximum Likelihood Estimates

			Standard	Wald	
Parameter	DF	Estimate	Error	Chi-Square	Pr > ChiSq
Intercept	1	9.7773	35.4872	0.0759	0.7829
drug	1	-19.2725	40.2338	0.2295	0.6319

The LOGISTIC Procedure

WARNING: The validity of the model fit is questionable.

Odds Ratio Estimates

	Point 95		s Wald
Effect	Estimate	Confide	ence Limits
drua	<0.001	<0.001	>999.999

Association of Predicted Probabilities and Observed Responses

Percent	Concordant	100.0	Somers' D	1.000
Percent	Discordant	0.0	Gamma	1.000
Percent	Tied	0.0	Tau-a	0.406
Pairs		518	С	1.000

Wald Confidence Interval for Parameters

95% Confidence Limits

79.3309

59.5842

-59.7764

-98.1293

Estimate

9.7773

-19.2725



The LOGISTIC Procedure

Exact Conditional Analysis

Conditional Exact Tests

			p-Va	alue
Effect	Test	Statistic	Exact	Mid
Intercept	Score	14.0000	0.0001	<.0001
	Probability	0.000061	0.0001	<.0001
drug	Score	50.0000	<.0001	<.0001
	Probability	7.74E-13	<.0001	<.0001

Exact Parameter Estimates

		95% Con	fidence		
Parameter	Estimate	imate Limits		p-Value	
Intercept	2.9807*	1.1991	Infinity	0.0001	
drug	-6.4343*	-Infinity	-4.1539	<.0001	

NOTE: * indicates a median unbiased estimate.

The SAS output for asymptotic and exact inference when considering the Quasi-Complete Separation data set is presented below.

Quasi-Complete Separation

	Mutation=NO	Mutation=YES	Total
Drug=NON-EXPOSURE	12	0	12
Drug=EXPOSURE	25	14	39
Total	37	14	51

The LOGISTIC Procedure

 ${\bf Model\ Information}$

Data Set

Response Variable

Number of Response Levels

Number of Observations

Model

Optimization Technique

WORK.TEST

mutation

State

Motal

Work.TEST

Mutation

State

Mutation

Mutation

State

WORK.TEST

Mutation

State

Mutation

State

Mutation

State

Mutation

State

State

WORK.TEST

Mutation

State

Mutation

State

Response Profile

Ordered		Total
Value	mutation	Frequency
1	1	14
2	0	37

Probability modeled is mutation=1.

Model Convergence Status

Quasi-complete separation of data points detected.

WARNING: The maximum likelihood estimate may not exist.

WARNING: The LOGISTIC procedure continues in spite of the above warning. Results shown are based on the last maximum likelihood iteration. Validity of the model fit is questionable.

Model Fit Statistics

			Intercept
		Intercept	and
C	riterion	Only	Covariates
A:	IC	61.945	54.920
S	C	63.877	58.784
- 2	2 Log L	59.945	50.920

Testing Global Null Hypothesis: BETA=0

Test	Chi-Square	DF	Pr > ChiSq
Likelihood Ratio	9.0242	1	0.0027
Score	5.9376	1	0.0148
Wald	0.0028	1	0.9581

Analysis of Maximum Likelihood Estimates

			Standard	Wald	
Parameter	DF	Estimate	Error	Chi-Square	Pr > ChiSq
Intercept	1	-13.4954	246.0	0.0030	0.9562
drug	1	12.9155	246.0	0.0028	0.9581

The LOGISTIC Procedure

WARNING: The validity of the model fit is questionable.

Odds Ratio Estimates

	Point	95% Wald
Effect	Estimate	Confidence Limits
dnua	>000 000	<0.001 >000.000

Association of Predicted Probabilities and Observed Responses

Percent	Concordant	32.4	Somers' D	0.324
Percent	Discordant	0.0	Gamma	1.000
Percent	Tied	67.6	Tau-a	0.132
Pairs		518	С	0.662

Wald Confidence Interval for Parameters

Parameter	Estimate	95% Confidence	Limits
Intercept	-13.4954	-495.6	468.6
drug	12.9155	-469.2	495.0

The LOGISTIC Procedure

Exact Conditional Analysis

Conditional Exact Tests

			p-Va	alue
Effect	Test	Statistic	Exact	Mid
Intercept	Score	12.0000	0.0005	0.0004
220. 000				
	Probability	0.000244	0.0005	0.0004
drug	Score	5.8212	0.0224	0.0166
	Probability	0.0117	0.0224	0.0166
	-			

Exact Parameter Estimates

		95% Con	ridence		
Parameter	Estimate	Limits		p-Value	
Intercept	-2.8224*	-Infinity	-1.0219	0.0005	
drug	2.1708*	0.2484	Infinity	0.0233	

NOTE: * indicates a median unbiased estimate.

The SAS output for asymptotic and exact inference when considering the Overlap data set is presented below.

Overlap

	Mutation=NO	Mutation=YES	Total
Drug=NON-EXPOSURE	9	3	12
Drug=EXPOSURE	25	14	39
Total	34	17	51

The LOGISTIC Procedure

Model Information

Data Set	WORK.TEST
Response Variable	mutation
Number of Response Levels	2
Number of Observations	51
Model	binary logit
Optimization Technique	Fisher's scoring

Response Profile

Total		Ordered
Frequency	mutation	Value
17	1	1
34	0	2

Probability modeled is mutation=1.

Model Convergence Status

Convergence criterion (GCONV=1E-8) satisfied.

Model Fit Statistics

Criterion	Intercept Only	Intercept and Covariates
AIC	66.924	68.416
SC	68.856	72.280
-2 Log L	64.924	64.416



Testing Global Null Hypothesis: BETA=0

Test	Chi-Square	DF	Pr > ChiSq
Likelihood Ratio	0.5080	1	0.4760
Score	0.4904	1	0.4838
Wald	0.4838	1	0.4867

Analysis of Maximum Likelihood Estimates

Parameter	DF	Estimate	Standard Error	Wald Chi-Square	Pr > ChiSq
Intercept	1	-1.0984	0.6666	2.7148	0.0994
drug		0.5186	0.7455	0.4838	0.4867

Odds Ratio Estimates

	Point	95% Wald
Effect	Estimate	Confidence Limits
drua	1,680	0.390 7.241

The LOGISTIC Procedure

Association of Predicted Probabilities and Observed Responses

Percent	Concordant	21.8	Somers' D	0.088
Percent	Discordant	13.0	Gamma	0.254
Percent	Tied	65.2	Tau-a	0.040
Pairs		578	С	0.544

Wald Confidence Interval for Parameters

Parameter	Estimate	95% Confidence	e Limits
Intercept	-1.0984	-2.4050	0.2082
drug	0.5186	-0.9427	1.9798

The LOGISTIC Procedure

Exact Conditional Analysis

Conditional Exact Tests

				p-Va	alue
	Effect	Test	Statistic	Exact	Mid
Collection of Biosta	Intercept	Score	3.0000	0.1460	0.1191
Collection of Biosta		Probability	0.0537	0.1460	0.1191
Research Archiv	drug	Score	0.4808	0.7278	0.6155
110300101171101111		Probability	0.2247	0.7278	0.6155

Exact Parameter Estimates

		95% Conf	idence	
Parameter	Estimate	Limi	p-Value	
Intercept	-1.0986	-2.8465	0.2894	0.1460
drug	0.5091	-1.0858	2.4087	0.7422

The SAS output for asymptotic and exact inference when considering the Linear Dependency data set is presented below.

Linear Dependency

	Mutation=NO	Mutation=YES	Total
Drug=NON-EXPOSURE	0	0	0
Drug=EXPOSURE	37	14	51
Total	37	14	51

The LOGISTIC Procedure

Model Information

Data Set	WORK.TEST
Response Variable	mutation
Number of Response Levels	2
Number of Observations	51
Model	binary logit
Optimization Technique	Fisher's scoring

Response Profile

Ordered		Total
Value	mutation	Frequency
1	1	14
2	0	37

Probability modeled is mutation=1.

Model Convergence Status

Convergence criterion (GCONV=1E-8) satisfied.

NOTE: The following parameters have been set to 0, since the variables are a linear combination of other variables as shown.

drug = Intercept

Analysis of Maximum Likelihood Estimates

Parameter	DF	Estimate	Standard Error	Wald Chi-Square	Pr > ChiSq
Intercept	1	-0.9719	0.3138	9.5933	0.0020
drug	Ü	Ü			

Wald Confidence Interval for Parameters

Parameter	Estimate	95% Confide	nce Limits
Intercept	-0.9719	-1.5869	-0.3569

There was no SAS Output for Exact Conditional Analysis.

Results

Complete Separation Data Set:

Unconditional maximum likelihood estimation (asymptotic inference)

The SAS output provided the following information:

Model Convergence Status

Complete separation of data points detected.

```
WARNING: The maximum likelihood estimate does not exist.
WARNING: The LOGISTIC procedure continues in spite of the above warning. Results shown are based on the last maximum likelihood iteration. Validity of the model fit is questionable.
```

The standard errors of the point estimates for the intercept (se= 35.4872) and drug (se= 40.2338) were large compared to the point estimates for the intercept (β = 9.7773) and drug (β = -19.2725). This is typically seen when the maximum likelihood parameter estimates do not converge during the modeling procedure. In addition, one sees that the p-values were both non-significant (p>0.05) for the intercept (p=0.7829) and drug (p=0.6319).

• Exact conditional estimation (exact inference)

The SAS output provided the following information:

A median unbiased estimate of the intercept (MU β = 2.9807, 95% exact CI [1.1991, ∞]) and drug (MU β = -6.4343, 95% exact CI [- ∞ , -4.1539]) are provided. One also sees that the exact p-values were both significant (intercept, p=0.0001; drug, p<0.0001).

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Quasi-Complete Separation Data Set:

• Unconditional maximum likelihood estimation (asymptotic inference)

The SAS output provided the following information:

```
Model Convergence Status
```

Quasi-complete separation of data points detected.

```
WARNING: The maximum likelihood estimate may not exist.
WARNING: The LOGISTIC procedure continues in spite of the above warning. Results shown are based on the last maximum likelihood iteration. Validity of the model fit is questionable.
```

The standard errors of the point estimates for the intercept (se= 246.0) and drug (se= 246.0) were very large compared to the point estimates for the intercept (β = -13.4954) and drug (β = 12.9155). This is typically seen when the maximum likelihood parameter estimates do not converge during the modeling procedure. In addition, one sees that the p-values were both non-significant (p>0.05) for the intercept (p=0.9562) and drug (p=0.9581).

• Exact conditional estimation (exact inference)

The SAS output provided the following information:

A median unbiased estimate of the intercept (MU β = -2.8224, 95% exact CI [- ∞ , -1.0219]) and drug (MU β = 2.1708, 95% exact CI [0.2484, ∞]) are provided. One also sees that the exact p-values were both significant (intercept, p=0.0005; drug, p<0.0233).

Overlap Data Set:

Unconditional maximum likelihood estimation (asymptotic inference)

The SAS output provided the following information:

```
Model Convergence Status

Convergence criterion (GCONV=1E-8) satisfied.
```

The standard errors of the point estimates for the intercept (se= 0.6666) and drug (se= 0.7455) were of reasonable size compared to the point estimates for the intercept (β = -1.0984) and drug (β = 0.5186). This is typically seen when the maximum likelihood parameter estimates does converge during the modeling procedure. In addition, one sees that the p-values were both non-significant (p>0.05) for the intercept (p=0.0994) and drug (p=0.4867).

• Exact conditional estimation (exact inference)

The SAS output provided the following information:

An exact estimate of the intercept (β = -1.0986, 95% exact CI [-2.8465, 0.2894]) and drug (β = 0.5091, 95% exact CI [-1.0858, 2.4087]) are provided. One also sees that the exact p-values were both non-significant (intercept, p=0.1460; drug, p=0.7422).

Linear Dependency Data Set:

• Unconditional maximum likelihood estimation (asymptotic inference)

The SAS output provided the following information:

Model Convergence Status

Convergence criterion (GCONV=1E-8) satisfied.

NOTE: The following parameters have been set to 0, since the variables are a linear combination of other variables as shown.

```
drug = Intercept
```

The standard error of the point estimate for the intercept (se= 0.3138) was of reasonable size compared to the point estimates for the intercept (β = -0.9719). This is typically seen when the maximum likelihood parameter estimates does converge during the modeling procedure. The point estimate and the standard error of the point estimate for drug were not computed because of the linear dependency between the two parameters. The p-value for the intercept was significant (p=0.0020).

• Exact conditional estimation (exact inference)

Because of the linear dependency, no SAS output was generated for the exact conditional analysis.

Take Home Points

- The Maximum Likelihood Estimates **do not exist** when you have a data set with **complete separation**.
- The Maximum Likelihood Estimates **may not exist** when you have a data set with **quasi-complete separation**.
- The Maximum Likelihood Estimates **do exist** when you have a data set with **overlap**.
- The Maximum Likelihood Estimates **do exist** when you have a data set with **linear dependency**.
- Exact Conditional Logistic Regression Analysis can be more informative than Unconditional Maximum Likelihood Logistic Regression Analysis.

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